

估计量的评选标准: 无偏性, 有效性, 相合性

1. 无偏性: $E(\hat{\theta}) = \theta$

例1. $\bar{x}$ 是总体均值 μ 的无偏估计.

证明: $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$, $E(\bar{x}) = E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n} \sum_{i=1}^n \underbrace{E(x_i)}_{\mu} = \frac{1}{n} \cdot \sum_{i=1}^n \mu = \mu$

例2. 样本方差 $S^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$ 是总体方差 σ^2 的无偏估计

证明: $S^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 = \frac{1}{n-1} (\sum x_i^2 - n\bar{x}^2)$

$$\begin{aligned} \Rightarrow E(S^2) &= \frac{1}{n-1} \left(\sum_{i=1}^n E(x_i^2) - n E(\bar{x}^2) \right) \\ &= \frac{1}{n-1} \left(\sum_{i=1}^n (\sigma^2 + \mu^2) - n \cdot (D(\bar{x}) + (E(\bar{x}))^2) \right) \\ &= \frac{1}{n-1} \left(n\sigma^2 + n\mu^2 - n \cdot \left(\frac{\sigma^2}{n} + \mu^2 \right) \right) \\ &= \frac{1}{n-1} (n-1)\sigma^2 = \sigma^2 \end{aligned}$$

$$\begin{aligned} \sum_{i=1}^n (x_i - \bar{x})^2 &= \sum_{i=1}^n (x_i^2 - 2x_i\bar{x} + \bar{x}^2) \\ &= \sum_{i=1}^n x_i^2 - 2\bar{x} \cdot \sum_{i=1}^n x_i + n\bar{x}^2 \\ &= \sum_{i=1}^n x_i^2 - 2 \cdot \frac{1}{n} \sum_{i=1}^n x_i \cdot \sum_{i=1}^n x_i + n\bar{x}^2 \\ &= \sum_{i=1}^n x_i^2 - 2n \cdot \left(\frac{1}{n} \sum_{i=1}^n x_i \right) \cdot \left(\frac{1}{n} \sum_{i=1}^n x_i \right) \\ &\quad + n\bar{x}^2 \\ &= \sum_{i=1}^n x_i^2 - n\bar{x}^2 \end{aligned}$$

$$\begin{aligned} D(X) &= E(X^2) - (E(X))^2 \\ \Rightarrow E(X^2) &= D(X) + (E(X))^2 \end{aligned}$$

2. 有效性: $D(\hat{\theta}_1) \leq D(\hat{\theta}_2)$ 称 $\hat{\theta}_1$ 比 $\hat{\theta}_2$ 更有效.

例3. 设 x_1, x_2, x_3 为来自于均值为 μ , 方差为 σ^2 的总体的一个样本. $\hat{\theta}_1 = \bar{x} = \frac{1}{3}(x_1 + x_2 + x_3)$

$\hat{\theta}_2 = \frac{1}{2}x_1 + \frac{1}{3}x_2 + \frac{1}{6}x_3$. 试证明 $\hat{\theta}_1$ 和 $\hat{\theta}_2$ 都是 μ 的无偏估计并确定哪一个估计量更有效

证明: $E(\hat{\theta}_1) = E\left(\frac{1}{3}(x_1 + x_2 + x_3)\right) = \frac{1}{3} (E(x_1) + E(x_2) + E(x_3)) = \frac{1}{3} \cdot 3\mu = \mu$

$E(\hat{\theta}_2) = E\left(\frac{1}{2}x_1 + \frac{1}{3}x_2 + \frac{1}{6}x_3\right) = \frac{1}{2}E(x_1) + \frac{1}{3}E(x_2) + \frac{1}{6}E(x_3) = \frac{1}{2}\mu + \frac{1}{3}\mu + \frac{1}{6}\mu = \mu$

$\Rightarrow \hat{\theta}_1$ 和 $\hat{\theta}_2$ 都是 μ 的无偏估计.

$$\begin{aligned} D(\hat{\theta}_1) &= D(\bar{x}) = D\left(\frac{1}{3}(x_1 + x_2 + x_3)\right) = \frac{1}{9} (Dx_1 + Dx_2 + Dx_3) = \frac{1}{9} \cdot (\sigma^2 + \sigma^2 + \sigma^2) \\ &= \frac{1}{3} \sigma^2 \end{aligned}$$

$$D(\hat{\theta}_2) = D\left(\frac{1}{2}\alpha_1 + \frac{1}{3}\alpha_2 + \frac{1}{6}\alpha_3\right) = \frac{1}{4}\sigma^2 + \frac{1}{9}\sigma^2 + \frac{1}{36}\sigma^2 = \frac{14}{36}\sigma^2$$

$$\therefore D(\hat{\theta}_1) < D(\hat{\theta}_2) \Rightarrow \hat{\theta}_1 \text{ 比 } \hat{\theta}_2 \text{ 更有效.}$$

3. 相容性 (一致性): $\lim_{n \rightarrow \infty} P(|\hat{\theta} - \theta| < \varepsilon) = 1$ ($\hat{\theta}$ 依概率收敛于 θ)

例 4. $\bar{x}$ 是总体均值 μ 的相容估计.

证明: (切比雪夫大数定律: $P\{|\bar{x} - \mu| < \varepsilon\} \geq 1 - \frac{D\bar{x}}{\varepsilon^2}$ $P\{|\bar{x} - E(\bar{x})| < \varepsilon\} \geq 1 - \frac{D\bar{x}}{\varepsilon^2}$)

$$P\{|\bar{x} - \mu| > \varepsilon\} \leq \frac{D\bar{x}}{\varepsilon^2}.$$

$$P\{|\bar{x} - E(\bar{x})| < \varepsilon\} \geq 1 - \frac{D\bar{x}}{\varepsilon^2} \quad \because E(\bar{x}) = \mu$$

$$\Rightarrow P\{|\bar{x} - \mu| < \varepsilon\} \geq 1 - \frac{D\bar{x}}{\varepsilon^2} = 1 - \frac{1}{\varepsilon^2} \cdot \frac{\sigma^2}{n}$$

$$\Rightarrow \lim_{n \rightarrow \infty} P\{|\bar{x} - \mu| < \varepsilon\} \geq 1 \Rightarrow \lim_{n \rightarrow \infty} P\{|\bar{x} - \mu| < \varepsilon\} = 1$$

定理: 若 $\lim_{n \rightarrow \infty} E(\hat{\theta}) = \theta$, $\lim_{n \rightarrow \infty} D(\hat{\theta}) = 0 \Rightarrow \hat{\theta}$ 是 θ 的相容性估计.

$$\begin{cases} D(\bar{x}) = \frac{\sigma^2}{n} \\ E(\bar{x}) = \mu \end{cases} \quad \lim_{n \rightarrow \infty} D(\bar{x}) = 0 \quad \therefore \bar{x} \text{ 是 } \mu \text{ 的相容性估计.}$$